Variational Inference of overparameterized Bayesian Neural Networks: a theoretical and empirical study

Anna Korba CREST, ENSAE, Institut Polytechnique de Paris

Laplace demon seminar

Joint work with Tom Huix, Szymon Majewski, Eric Moulines (CMAP, Polytechnique) and Alain Durmus (ENS Cachan).



Problem and Motivation

VI for BNN

Identifying well-posed regimes for the ELBO with product priors

Experiments

Sampling

Problem: Sample (=generate new examples) from a target distribution π over \mathbb{R}^d , whose density w.r.t. Lebesgue measure is known up to an intractable normalisation constant Z:

$$\pi(w) = \frac{\tilde{\pi}(w)}{Z}, \quad \tilde{\pi} \text{ known, } Z \text{ unknown.}$$

Main application: Bayesian inference, where π is the posterior distribution over parameters of a model.

Let $\mathcal{D} = (x_i, y_i)_{i=1}^m$ a dataset of labelled examples $(x_i, y_i) \stackrel{i.i.d.}{\sim} P_{data}$. Assume an underlying model parametrized by *w*, e.g. :

$$y = g(x, w) + \epsilon, \quad \epsilon \sim \mathcal{N}(0, I)$$

Goal: learn the best distribution over *w* to fit the data.

Let $\mathcal{D} = (x_i, y_i)_{i=1}^m$ a dataset of labelled examples $(x_i, y_i) \stackrel{i.i.d.}{\sim} P_{data}$. Assume an underlying model parametrized by *w*, e.g. :

$$y = g(x, w) + \epsilon, \quad \epsilon \sim \mathcal{N}(0, I)$$

Goal: learn the best distribution over *w* to fit the data.

1. Compute the Likelihood:

$$p(\mathcal{D}|\boldsymbol{w}) = \prod_{i=1}^{m} p(y_i|\boldsymbol{w}, x_i) \propto \exp\left(-\frac{1}{2}\sum_{i=1}^{m} \|y_i - g(x_i, \boldsymbol{w})\|^2\right).$$

Let $\mathcal{D} = (x_i, y_i)_{i=1}^m$ a dataset of labelled examples $(x_i, y_i) \stackrel{i.i.d.}{\sim} P_{data}$. Assume an underlying model parametrized by *w*, e.g. :

$$y = g(x, w) + \epsilon, \quad \epsilon \sim \mathcal{N}(0, I)$$

Goal: learn the best distribution over *w* to fit the data.

1. Compute the Likelihood:

$$p(\mathcal{D}|w) = \prod_{i=1}^m p(y_i|w, x_i) \propto \exp\left(-\frac{1}{2}\sum_{i=1}^m \|y_i - g(x_i, w)\|^2\right).$$

2. Choose a prior distribution on the parameter:

$$w \sim p$$
, e.g. $p(w) \propto \exp\left(-\frac{\|w\|^2}{2}\right)$.

Let $\mathcal{D} = (x_i, y_i)_{i=1}^m$ a dataset of labelled examples $(x_i, y_i) \stackrel{i.i.d.}{\sim} P_{data}$. Assume an underlying model parametrized by *w*, e.g. :

$$y = g(x, w) + \epsilon, \quad \epsilon \sim \mathcal{N}(0, I)$$

Goal: learn the best distribution over *w* to fit the data.

1. Compute the Likelihood:

$$p(\mathcal{D}|\boldsymbol{w}) = \prod_{i=1}^{m} p(y_i|\boldsymbol{w}, x_i) \propto \exp\left(-\frac{1}{2}\sum_{i=1}^{m} \|y_i - g(x_i, \boldsymbol{w})\|^2\right).$$

2. Choose a prior distribution on the parameter:

$$w \sim p$$
, e.g. $p(w) \propto \exp\left(-\frac{\|w\|^2}{2}\right)$.

3. Bayes' rule yields:

$$\pi(w) := p(w|\mathcal{D}) = \frac{p(\mathcal{D}|w)p(w)}{Z} \quad Z = \int_{\mathbb{R}^d} p(\mathcal{D}|w)p(w)dw$$

i.e.
$$\pi(w) \propto \exp(-V(w))$$
, $V(w) = \frac{1}{2} \sum_{i=1}^{m} \|y_i - g(x_i, w)\|^2 + \frac{\|w\|^2}{2}$.

- π is needed both for
 - prediction for a new input *x*: $y_{pred} = \int_{\mathbb{R}^d} g(x, w) d\pi(w)$

measure uncertainty on the prediction.

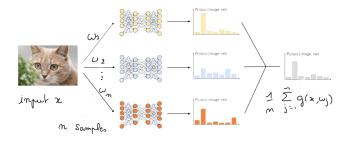
 π is needed both for

• prediction for a new input *x*: $y_{pred} = \int_{\mathbb{R}^d} g(x, w) d\pi(w)$

measure uncertainty on the prediction.

Given a discrete approximation $\mu_n = \frac{1}{n} \sum_{i=1}^n \delta_{w_i}$ of π :

$$y_{pred} \approx rac{1}{n} \sum_{j=1}^n g(x, w_j).$$



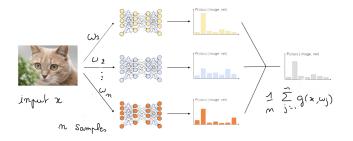
 π is needed both for

• prediction for a new input *x*: $y_{pred} = \int_{\mathbb{R}^d} g(x, w) d\pi(w)$

measure uncertainty on the prediction.

Given a discrete approximation $\mu_n = \frac{1}{n} \sum_{i=1}^n \delta_{w_i}$ of π :

$$y_{pred} \approx rac{1}{n} \sum_{j=1}^{n} g(x, w_j).$$



Question: how can we approximate π ?

Main methods for sampling

Markov Chain Monte Carlo Methods (MCMC) generate a Markov chain whose law converges to π ∝ exp(−V)

Example: Langevin Monte Carlo (LMC)

$$w_{l+1} = w_l - \gamma \nabla V(w_l) + \sqrt{2\gamma} \epsilon_l, \ \epsilon_l \sim \mathcal{N}(0, I_d)$$

other example: Hamiltonian Monte Carlo

 Variational inference (VI) methods approximate π with a parametric distribution by solving

$$\min_{\theta \in \Theta} \mathsf{KL}(\boldsymbol{p}_{\theta}|\pi)$$

Difficult cases : non-convex potentials

Recall that

$$\pi(w) \propto \exp\left(-V(w)\right), \quad V(w) = \underbrace{\sum_{i=1}^{m} \|y_i - g(x_i, w)\|^2}_{\text{loss}} + \frac{\|w\|^2}{2}.$$

- ► if V is convex (e.g. g(x, w) = ⟨w, x⟩) many sampling MCMC methods come with theoretical guarantees,
- but if its not (e.g. g(x, w) is a neural network), the situation is much more delicate



A highly nonconvex neural net loss surface. From https://www.telesens.co/2019/01/16/neural-network-loss-visualization. MCMC methods do not scale and require too many iterations (≈ 10⁴) see [Izmailov et al., 2021] that run HMC over 512 Tensor processing unit (TPU) devices to obtain baselines on CIFAR10

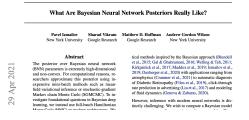


Figure: Long oral ICML 2021.

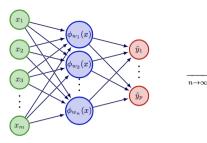
VI remains a standard approach in Bayesian Deep Learning

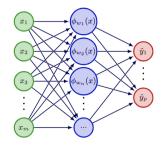
Question: What can we say on the validity or limitations of VI for Bayesian Neural Networks (BNN)?

especially in the current, **overparametrized** regime era for neural networks.

Infinite width neural network

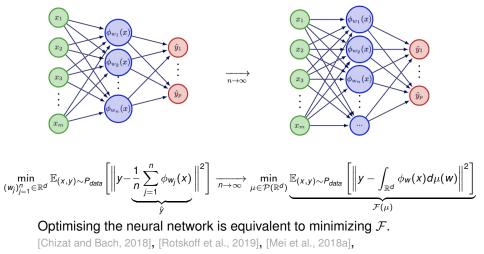
consider a one-hidden-layer neural network, denote $\phi_{w_j}(x) = a_j \sigma(\langle b_j, x \rangle)$ the output of neuron *j*.





Infinite width neural network

consider a one-hidden-layer neural network, denote $\phi_{w_j}(x) = a_j \sigma(\langle b_j, x \rangle)$ the output of neuron *j*.



[Arbel et al., 2019]...

Idea: consider a similar regime for VI on BNN.



Problem and Motivation

VI for BNN

Identifying well-posed regimes for the ELBO with product priors

Experiments

Assume we have access to $\{(x_i, y_i)\}_{i=1}^p$ samples from the data distribution on X × Y.

for each input $x \in X$, the output prediction $f_{W} : X \to \mathbb{R}^{d_{Y}}$ of the neural network can be written as:

$$f_{w}(x) = \frac{1}{N} \sum_{j=1}^{N} s(w_j, x), \text{ with } s(w_j, x) = a_j \sigma(\langle b_j, x \rangle),$$

 $w_j = (a_j, b_j) \in \mathbb{R}^d, \ w \in \mathbb{R}^{N imes d}.$

Assume we have access to $\{(x_i, y_i)\}_{i=1}^p$ samples from the data distribution on X × Y.

for each input $x \in X$, the output prediction $f_{w} : X \to \mathbb{R}^{d_{Y}}$ of the neural network can be written as:

$$f_{\boldsymbol{w}}(x) = \frac{1}{N} \sum_{j=1}^{N} s(\boldsymbol{w}_j, x), \text{ with } s(\boldsymbol{w}_j, x) = a_j \sigma(\langle \boldsymbol{b}_j, x \rangle),$$

 $\boldsymbol{w}_j = (a_j, b_j) \in \mathbb{R}^d, \ \boldsymbol{w} \in \mathbb{R}^{N \times d}.$

Given a loss function $\ell: Y \times Y \to \mathbb{R}_+$, the likelihood is defined as

$$L(y|x, \mathbf{W}) \propto \exp(-\ell(f_{\mathbf{W}}(x), y))$$
.

Assume we have access to $\{(x_i, y_i)\}_{i=1}^p$ samples from the data distribution on X × Y.

for each input $x \in X$, the output prediction $f_{w} : X \to \mathbb{R}^{d_{Y}}$ of the neural network can be written as:

$$f_{\boldsymbol{w}}(x) = \frac{1}{N} \sum_{j=1}^{N} s(w_j, x), \text{ with } s(w_j, x) = a_j \sigma(\langle b_j, x \rangle),$$

 $w_j = (a_j, b_j) \in \mathbb{R}^d, \ \boldsymbol{w} \in \mathbb{R}^{N \times d}.$

Given a loss function $\ell: Y \times Y \to \mathbb{R}_+,$ the likelihood is defined as

$$L(y|x, \mathbf{W}) \propto \exp(-\ell(f_{\mathbf{W}}(x), y))$$
.

Then, choosing a prior pdf P_0 on \boldsymbol{w} , the posterior pdf P of the weights is

$$P(\boldsymbol{w}) = \frac{P_0(\boldsymbol{w}) \prod_{i=1}^p L(y_i | x_i, \boldsymbol{w})}{Z}$$

Recall that VI considers a variational family of pdfs $\mathscr{F}_{\Theta} = \{q_{\theta} : \theta \in \Theta\}$ and solves

$$\theta^* \in \operatorname{argmin}_{\theta \in \Theta} \mathsf{KL}(q_{\theta} | P), \quad P(w) = \frac{P_0(w) \prod_{i=1}^{p} \mathrm{L}(y_i | x_i, w)}{Z}.$$

Recall that VI considers a variational family of pdfs $\mathscr{F}_{\Theta} = \{q_{\theta} : \theta \in \Theta\}$ and solves

$$\theta^* \in \operatorname{argmin}_{\theta \in \Theta} \mathsf{KL}(q_{\theta} | P), \quad P(w) = \frac{P_0(w) \prod_{i=1}^p \mathrm{L}(y_i | x_i, w)}{Z}.$$

It is equivalent to maximizing the Evidence Lower Bound (ELBO) defined for any $\theta \in \Theta$ by:

$$\text{ELBO}^{N}(\boldsymbol{\theta}) = \underbrace{-\text{KL}(\boldsymbol{q}_{\boldsymbol{\theta}} \mid \boldsymbol{P}_{0})}_{(1) \text{ penalty term}} + \underbrace{\sum_{i=1}^{p} \int_{\mathbb{R}^{N \times d}} \log L(\boldsymbol{y}_{i} \mid \boldsymbol{x}_{i}, \boldsymbol{w}) \boldsymbol{q}_{\boldsymbol{\theta}}(\boldsymbol{w}) d\boldsymbol{w}}_{(2) \text{ data fitting term}}.$$

Recall that VI considers a variational family of pdfs $\mathscr{F}_{\Theta} = \{q_{\theta} : \theta \in \Theta\}$ and solves

$$\theta^* \in \operatorname{argmin}_{\theta \in \Theta} \mathsf{KL}(q_{\theta} | P), \quad P(w) = \frac{P_0(w) \prod_{i=1}^p \mathrm{L}(y_i | x_i, w)}{Z}.$$

It is equivalent to maximizing the Evidence Lower Bound (ELBO) defined for any $\theta \in \Theta$ by:

$$\text{ELBO}^{N}(\boldsymbol{\theta}) = \underbrace{-\text{KL}(\boldsymbol{q}_{\boldsymbol{\theta}} \mid \boldsymbol{P}_{0})}_{(1) \text{ penalty term}} + \underbrace{\sum_{i=1}^{p} \int_{\mathbb{R}^{N \times d}} \log L(\boldsymbol{y}_{i} \mid \boldsymbol{x}_{i}, \boldsymbol{w}) \boldsymbol{q}_{\boldsymbol{\theta}}(\boldsymbol{w}) d\boldsymbol{w}}_{(2) \text{ data fitting term}}$$

In practice, it is common to consider a tempered ELBO^N:

[Zhang et al., 2018, Khan et al., 2018, Osawa et al., 2019, Ashukha et al., 2020]

$$\mathrm{ELBO}^{N}_{\eta}(heta) = -\eta \, \mathrm{KL}(q_{ heta} \mid P_{0}) + \sum_{i=1}^{p} \int_{\mathbb{R}^{N imes d}} \log \mathrm{L}(y_{i} \mid x_{i}, oldsymbol{w}) \mathrm{d}oldsymbol{w} \; .$$

 $\mathrm{ELBO}_{\eta}^{N} \iff \mathrm{ELBO}^{N}$ where *P* is replaced by a tempered posterior $P_{T_{N}} \propto \mathrm{L}^{1/\eta} P_{0}$ [Wenzel et al., 2020, Wilson and Izmailov, 2020].

In the VI literature, one can find for instance:

Reference	temperature η_N
[Zhang et al., 2018]	$\eta \in \{1/2, \dots, 1/10\}$
[Osawa et al., 2019]	$\eta \in \{1/5,\ldots,1/10\}$
[Ashukha et al., 2020]	η from 10 ⁻⁵ to 10 ⁻³

 η reweights the KL term and is typically smaller than 1 on current prediction tasks/neural nets architecture. From:

How Good is the Bayes Posterior in Deep Neural Networks Really?

Florian Wenzel^{*1} Kevin Roth^{*+2} Bastiaan S. Veeling^{*+31} Jakub Świątkowski⁴⁺ Linh Tran⁵⁺ Stephan Mandt⁶⁺ Jasper Snoek¹ Tim Salimans¹ Rodolphe Jenatton¹ Sebastian Nowzin⁷⁺



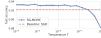


Figure 1. The "cold posterior" effect: for a ResNet-20 on CIFAR-10 we can improve the generalization performance significantly by cooling the posterior with a temperature $T \ll 1$, deviating from the Bayes posterior $p(\theta|D) \propto \exp(-U(\theta)/T)$ at T = 1.

Figure: Cold posteriors for training BNN with stochastic gradient Stochastic Gradient Markov chain Monte Carlo methods. Long oral ICML 2020.

Informally: Why tempering?

Idea: in parametric approaches, the model capacity (which is determined by the number of neurons and the neural network architecture) is chosen by the user; hence it may be mispecified.

Informally: Why tempering?

Idea: in parametric approaches, the model capacity (which is determined by the number of neurons and the neural network architecture) is chosen by the user; hence it may be mispecified.

It has been shown that tempered models may have better statistical properties than non tempered ones, e.g. for Generalized Linear Models

[Grünwald, 2012, Grünwald and Van Ommen, 2017, Bhattacharya et al., 2019, Heide et al., 2020, Grunwald et al., 2021] - not clear how this extends to BNN.

Informally: Why tempering?

Idea: in parametric approaches, the model capacity (which is determined by the number of neurons and the neural network architecture) is chosen by the user; hence it may be mispecified.

It has been shown that tempered models may have better statistical properties than non tempered ones, e.g. for Generalized Linear Models

[Grünwald, 2012, Grünwald and Van Ommen, 2017, Bhattacharya et al., 2019, Heide et al., 2020, Grunwald et al., 2021] - not clear how this extends to BNN.

Our work: study the impact of the choice of the cooling parameter η_N in the overparametrized regime (1 hidden layer neural net).

Our model - independent neurons, diagonal Gaussians

We consider a prior on $\pmb{w} \in \mathbb{R}^{N \times d}$ which factorize over the weights, i.e., of the form

$$\mathsf{P}_0(\boldsymbol{w}) = \prod_{j=1}^N \mathsf{P}_0^1(w_j) \; ,$$

and similarly for the variational posterior

$$q_{ heta}(oldsymbol{w}) = \prod_{i=1}^N q_{ heta_i}^1(w_i)$$

where P_0^1 and $\{q_{\theta_i}^1\}_{i=1}^N$ are distributions over \mathbb{R}^d .

For each neuron, we consider $q^1_{ heta} = (T_{ heta})_{\#}\gamma$ where $\gamma = \mathcal{N}(\mathbf{0}, I_d)$ and

$$\mathbf{T}_{\theta}: \mathbf{Z} \mapsto \mu + \sigma \odot \mathbf{Z} , \quad \theta = (\mu, \sigma) \in \mathbb{R}^{2d}$$

where \odot is the component wise product.

In this case, $\theta = (\theta_1, \dots, \theta_N) \in \mathbb{R}^{N \times 2d}$.

Recall the tempered ELBO:

$$\mathrm{ELBO}_{\eta}^{N}(\boldsymbol{\theta}) = -\eta \, \mathrm{KL}(\boldsymbol{q}_{\boldsymbol{\theta}} \,|\, \boldsymbol{P}_{0}) + \sum_{i=1}^{p} \int_{\mathbb{R}^{N \times d}} \log \mathrm{L}(\boldsymbol{y}_{i} | \boldsymbol{x}_{i}, \boldsymbol{w}) \boldsymbol{q}_{\boldsymbol{\theta}}(\boldsymbol{w}) \mathrm{d}\boldsymbol{w} \;.$$

Recall the tempered ELBO:

$$\text{ELBO}_{\eta}^{N}(\boldsymbol{\theta}) = -\eta \, \text{KL}(\boldsymbol{q}_{\boldsymbol{\theta}} \mid \boldsymbol{P}_{0}) + \sum_{i=1}^{p} \int_{\mathbb{R}^{N \times d}} \log L(\boldsymbol{y}_{i} | \boldsymbol{x}_{i}, \boldsymbol{w}) \boldsymbol{q}_{\boldsymbol{\theta}}(\boldsymbol{w}) \mathrm{d}\boldsymbol{w} \; .$$

To make the dependence in *N* more explicit, we can rewrite it as:

$$\text{ELBO}_{\eta}^{N}(\boldsymbol{\theta}) = -\eta \sum_{j=1}^{N} \text{KL}(\boldsymbol{q}_{\theta_{j}}^{1} | \boldsymbol{P}_{0}^{1}) - \sum_{i=1}^{p} \text{G}_{\Theta}^{N}(\boldsymbol{\theta}; (\boldsymbol{x}_{i}, \boldsymbol{y}_{i}))$$
(2)

where, denoting the output of a neuron parametrized by $\theta \in \mathbb{R}^d$ for an input x_i by

$$\phi(\theta, z, x_i) = s(T_{\theta}(z), x_i) ,$$

and $\boldsymbol{z} = (z_1, \dots, z_N) \in \mathbb{R}^{d \times N},$
$$G_{\Theta}^{N}(\theta; (x, y)) = \int \ell \left(y, \sum_{j=1}^{N} \frac{\phi(\theta_j, z_j, x)}{N} \right) \gamma^{\otimes N}(d\boldsymbol{z}) .$$

Recall the tempered ELBO:

$$\text{ELBO}_{\eta}^{N}(\boldsymbol{\theta}) = -\eta \, \text{KL}(\boldsymbol{q}_{\boldsymbol{\theta}} \mid \boldsymbol{P}_{0}) + \sum_{i=1}^{p} \int_{\mathbb{R}^{N \times d}} \log L(\boldsymbol{y}_{i} \mid \boldsymbol{x}_{i}, \boldsymbol{w}) \boldsymbol{q}_{\boldsymbol{\theta}}(\boldsymbol{w}) \mathrm{d}\boldsymbol{w} \; .$$

To make the dependence in *N* more explicit, we can rewrite it as:

$$\text{ELBO}_{\eta}^{N}(\boldsymbol{\theta}) = -\eta \sum_{j=1}^{N} \text{KL}(\boldsymbol{q}_{\theta_{j}}^{1} | \boldsymbol{P}_{0}^{1}) - \sum_{i=1}^{p} \text{G}_{\Theta}^{N}(\boldsymbol{\theta}; (\boldsymbol{x}_{i}, \boldsymbol{y}_{i}))$$
(2)

where, denoting the output of a neuron parametrized by $\theta \in \mathbb{R}^d$ for an input x_i by

$$\phi(\theta, \boldsymbol{z}, \boldsymbol{x}_i) = \boldsymbol{s}(\mathrm{T}_{\theta}(\boldsymbol{z}), \boldsymbol{x}_i) \; ,$$

and $\boldsymbol{z} = (z_1, \ldots, z_N) \in \mathbb{R}^{d \times N}$,

$$\mathrm{G}^{N}_{\Theta}(oldsymbol{ heta};(x,y)) = \int \ell\left(y,\sum_{j=1}^{N}rac{\phi(heta_{j},z_{j},x)}{N}
ight)\gamma^{\otimes N}(\mathrm{d}oldsymbol{z}) \ .$$

Problem: (1) scales as $\mathcal{O}(N)$, while (2) scales as $\mathcal{O}(p)$ and does not grow with *N* if the variance of q_{θ} does not scale with *N*.

 \implies (1) becomes predominant as $N \rightarrow \infty$!

The ELBO in our model

Proposition. Let $\theta^{*,N} = \operatorname{argmax}_{\theta \in \Theta} \operatorname{ELBO}^{N}(\theta)$. Assume that $P_0 \in \mathscr{F}_{\Theta}$ where \mathscr{F}_{Θ} are diagonal Gaussians, that *I* is the square loss or cross-entropy, Lipschitz activation functions for the neural network, and that X is compact. Then,

 $\mathsf{KL}(q_{\theta^{*,N}}, P_0) \to 0 \text{ as } N \to \infty.$

The ELBO in our model

Proposition. Let $\theta^{*,N} = \operatorname{argmax}_{\theta \in \Theta} \operatorname{ELBO}^{N}(\theta)$. Assume that $P_0 \in \mathscr{F}_{\Theta}$ where \mathscr{F}_{Θ} are diagonal Gaussians, that *I* is the square loss or cross-entropy, Lipschitz activation functions for the neural network, and that X is compact. Then,

$$\mathsf{KL}(q_{\theta^{*,N}}, P_0) \to 0 \text{ as } N \to \infty.$$

inspired from [Coker et al., 2021] that show a similar result when I is the

square loss and activation functions are odd.

The ELBO in our model

Proposition. Let $\theta^{*,N} = \operatorname{argmax}_{\theta \in \Theta} \operatorname{ELBO}^{N}(\theta)$. Assume that $P_0 \in \mathscr{F}_{\Theta}$ where \mathscr{F}_{Θ} are diagonal Gaussians, that *I* is the square loss or cross-entropy, Lipschitz activation functions for the neural network, and that X is compact. Then,

$$\mathsf{KL}(q_{\theta^{*,N}}, P_0) \to 0 \text{ as } N \to \infty.$$

inspired from [Coker et al., 2021] that show a similar result when *I* is the square loss and activation functions are odd.

Idea of the proof: By the optimality of $\theta^{\star,N}$, we have:

$$-\mathsf{KL}(q_{\theta^{*,N}}|P_0) - \mathcal{L}(q_{\theta^{*,N}}) = \mathrm{ELBO}^N(\theta^*) \geq \mathrm{ELBO}^N(\theta_0) = -\mathcal{L}(P_0)$$

Hence,

$$\mathsf{KL}(q_{{\boldsymbol{\theta}}^{*,N}}|P_0) \leq \mathcal{L}(P_0) - \mathcal{L}(q_{{\boldsymbol{\theta}}^{*,N}}).$$

Then show that both terms on the r.h.s. have the same finite limit.

Example of the square loss: we have

$$\mathcal{L}(q_{\theta}^{N}) = \sum_{i=1}^{p} \mathbb{E}_{\mathbf{w} \sim q_{\theta}^{N}} \left[\|y_{i}\|^{2} + \|f_{\mathbf{w}}(x_{i})\|^{2} - 2\langle y_{i}, f_{\mathbf{w}}(x_{i})\rangle + \log(Z) \right]$$

we first obtain:

$$\lim_{N\to\infty}\mathcal{L}(q_{\theta_0}^N)=\sum_{i=1}^p\|y_i\|^2+\log Z.$$

Furthermore,

$$\mathsf{KL}(q^{\mathsf{N}}_{{\boldsymbol{ heta}}^*}|q^{\mathsf{N}}_{{\boldsymbol{ heta}}_0}) \leq \mathcal{L}(q^{\mathsf{N}}_{{\boldsymbol{ heta}}_0}),$$

hence the KL is bounded by C_{KL} . Then we have we have:

$$\mathbb{E}_{\boldsymbol{w} \sim q_{\boldsymbol{\theta}^*}^N}[f_{\boldsymbol{w}}(x)] \leq \frac{F(\mathsf{KL}(q_{\boldsymbol{\theta}^*}^N, q_{\boldsymbol{\theta}_0}^N), \mathsf{X}, d_{\mathsf{Y}})}{\sqrt{N}} \leq \frac{F(C_{\mathsf{KL}}, \mathsf{X}, d_{\mathsf{Y}})}{\sqrt{N}}$$
$$\mathbb{E}_{\boldsymbol{w} \sim q_{\boldsymbol{\theta}^*}^N}[\|f_{\boldsymbol{w}}(x)\|^2] \leq \frac{G(\mathsf{KL}(q_{\boldsymbol{\theta}^*}^N, q_{\boldsymbol{\theta}_0}^N), \mathsf{X}, d_{\mathsf{Y}})}{\sqrt{N}} \leq \frac{G(C_{\mathsf{KL}}, \mathsf{X}, d_{\mathsf{Y}})}{\sqrt{N}}$$

Hence, we obtain:

$$\lim_{N\to\infty}\mathcal{L}(q_{\theta^*}^N)=\sum_{i=1}^p\|y_i\|^2+\log Z.$$



Problem and Motivation

VI for BNN

Identifying well-posed regimes for the ELBO with product priors

Experiments

First step: generalize the definition of ELBO_{η}^{N} defined in over $\mathbb{R}^{N \times 2d}$ to probability measures ν on \mathbb{R}^{2d} .

First step: generalize the definition of ELBO_{η}^{N} defined in over $\mathbb{R}^{N \times 2d}$ to probability measures ν on \mathbb{R}^{2d} .

Recall that

$$\text{ELBO}_{\eta}^{N}(\boldsymbol{\theta}) = -\eta \sum_{j=1}^{N} \text{KL}(\boldsymbol{q}_{\theta_{j}}^{1} | \boldsymbol{P}_{0}^{1}) - \sum_{i=1}^{p} \text{G}_{\Theta}^{N}(\boldsymbol{\theta}; (\boldsymbol{x}_{i}, \boldsymbol{y}_{i}))$$

where, denoting $\boldsymbol{z} = (z_1, \ldots, z_N) \in \mathbb{R}^{d \times N}$,

$$\mathrm{G}^{\mathsf{N}}_{\Theta}(oldsymbol{ heta};(x,y)) = \int \ell\left(y,\sum_{j=1}^{\mathsf{N}}rac{\phi(heta_j,z_j,x)}{\mathsf{N}}
ight) \gamma^{\otimes \mathsf{N}}(\mathrm{d} oldsymbol{z}) \;.$$

First step: generalize the definition of ELBO_{η}^{N} defined in over $\mathbb{R}^{N \times 2d}$ to probability measures ν on \mathbb{R}^{2d} .

Recall that

$$\text{ELBO}_{\eta}^{N}(\boldsymbol{\theta}) = -\eta \sum_{j=1}^{N} \text{KL}(\boldsymbol{q}_{\theta_{j}}^{1} | \boldsymbol{P}_{0}^{1}) - \sum_{i=1}^{p} \text{G}_{\Theta}^{N}(\boldsymbol{\theta}; (\boldsymbol{x}_{i}, \boldsymbol{y}_{i}))$$

where, denoting $\boldsymbol{z} = (z_1, \ldots, z_N) \in \mathbb{R}^{d \times N}$,

$$\mathrm{G}^{N}_{\Theta}(m{ heta};(x,y)) = \int \ell\left(y,\sum_{j=1}^{N}rac{\phi(heta_{j},z_{j},x)}{N}
ight)\gamma^{\otimes N}(\mathrm{d} m{z}) \;.$$

Define

$$\nu_N^{\theta} = \frac{1}{N} \sum_{i=1}^N \delta_{\theta_i} , \qquad (1)$$

Proposition For any $N \in \mathbb{N}$, there exists a function F_{η}^{N} defined over measures of the form (1), such that $\text{ELBO}_{\eta}^{N}(\theta) = F_{\eta}^{N}(\nu_{N}^{\theta})$ for any $\theta \in \mathbb{R}^{N \times 2d}$.

First step: generalize the definition of ELBO_{η}^{N} defined in over $\mathbb{R}^{N \times 2d}$ to probability measures ν on \mathbb{R}^{2d} .

Recall that

$$\text{ELBO}_{\eta}^{N}(\boldsymbol{\theta}) = -\eta \sum_{j=1}^{N} \text{KL}(\boldsymbol{q}_{\theta_{j}}^{1} | \boldsymbol{P}_{0}^{1}) - \sum_{i=1}^{p} \text{G}_{\Theta}^{N}(\boldsymbol{\theta}; (\boldsymbol{x}_{i}, \boldsymbol{y}_{i}))$$

where, denoting $\boldsymbol{z} = (z_1, \ldots, z_N) \in \mathbb{R}^{d \times N}$,

$$\mathrm{G}^{N}_{\Theta}(m{ heta};(x,y)) = \int \ell\left(y,\sum_{j=1}^{N}rac{\phi(heta_{j},z_{j},x)}{N}
ight)\gamma^{\otimes N}(\mathrm{d} m{z}) \;.$$

Define

$$\nu_N^{\theta} = \frac{1}{N} \sum_{i=1}^N \delta_{\theta_i} , \qquad (1)$$

Proposition For any $N \in \mathbb{N}$, there exists a function F_{η}^{N} defined over measures of the form (1), such that $\text{ELBO}_{\eta}^{N}(\theta) = F_{\eta}^{N}(\nu_{N}^{\theta})$ for any $\theta \in \mathbb{R}^{N \times 2d}$.

Problem: F_{η}^{N} cannot be non-trivially extended to a functional defined for a general probability measure on \mathbb{R}^{2d} .

We show that, when restricted to empirical probabilities, F_{η}^{N} is a perturbation, as $N \to +\infty$, of the functional \tilde{F}_{η}^{N} defined over all $\mathcal{P}(\mathbb{R}^{2d})$ by

$$\tilde{F}_{\eta}^{N}(\nu) = -\sum_{i=1}^{\rho} \tilde{G}(\nu; (\mathbf{x}_{i}, \mathbf{y}_{i})) - \eta \mathbf{N} \int \mathsf{KL}(\mathbf{q}_{\theta}^{1} | \mathbf{P}_{0}^{1}) \mathrm{d}\nu(\theta) ,$$

where

$$\tilde{G}(\nu; (\mathbf{x}, \mathbf{y})) = \ell \left(\mathbf{y}, \underbrace{\iint \phi(\theta, \mathbf{z}, \mathbf{x}) d\nu(\theta) d\gamma(\mathbf{z})}_{\iint \mathbf{s}(T_{\theta}(\mathbf{z}), \mathbf{x}) d\gamma(\mathbf{z}) d\nu(\theta)} \right) ,$$

Remark:

- \tilde{G} differs from G_{Θ}^{N} through the integration "inside" the loss

Theorem: Under mild assumptions on the loss, activation functions, prior, X, Y; there exists $C \ge 0$ such that for any $N, p \in \mathbb{N}$, $\{(x_i, y_i)\}_{i=1}^p \in (X \times Y)^p, \theta \in \Xi^N \text{ and } \eta > 0,$ $|\text{ELBO}_n^N(\theta) - \tilde{F}_n^N(\nu_N^\theta)| \le Cp/N$,

21/27

It is now much clearer how to define a **balanced functional** over $\mathcal{P}(\mathbb{R}^{2d})$.

We now set $\eta = \tau p / N$ with $\tau > 0$.

With this particular choice, \tilde{F}_{η}^{N} depends only on the number of observations *p* but no longer on the number of neurons *N*. We denote, for that particular choice of η_{N} ,

$$\mathcal{F}(\nu) = \rho^{-1} \tilde{\mathrm{F}}_{\eta}^{N}(\nu) = -\frac{1}{\rho} \sum_{i=1}^{\rho} \tilde{\mathrm{G}}(\nu; (x_i, y_i)) - \tau \int \mathrm{KL}(q_{\theta}^{1} | P_0^{1}) \mathrm{d}\nu(\theta) \ .$$



Problem and Motivation

VI for BNN

Identifying well-posed regimes for the ELBO with product priors

Experiments

We illustrate our findings and their practical implications for image classification on standard datasets (MNIST, CIFAR-10), with a simple one hidden layer architecture and a Resnet20 respectively.

We illustrate our findings and their practical implications for image classification on standard datasets (MNIST, CIFAR-10), with a simple one hidden layer architecture and a Resnet20 respectively.

For each neuron, we use a centered Gaussian prior with variance 1/5, following [Osawa et al., 2019]. We train each BNN by Bayes by Backprop [Blundell et al., 2015].

Metrics:

For an input $x \in X$, the predictive probability of a class c by a neural network with weights \boldsymbol{w} is defined by $\Psi_c(f_{\boldsymbol{w}}(x))$, where $\Psi_c(f_{\boldsymbol{w}}(x))$ denotes the c-th component of the softmax function applied to the output $f_{\boldsymbol{w}}(x) \in \mathbb{R}^{n_l}$ of the neural network.

- Accuracy: number of correct predictions
- ► NLL: $\sum_{i=1}^{p} \int_{\mathbb{R}^{N \times d}} \ell_{CE}(y_i, f_w(x_i)) q_\theta(w) dw$ where ℓ_{CE} is the cross-entropy loss
- ► ECE: measures if the predictive posterior is close to the true probability for each class c ∈ {1,..., n_l}.
- Confidence: conf(x) = max_{c∈{1,...,n_l}} Ψ_c(f_w(x)) averaged over all points x.

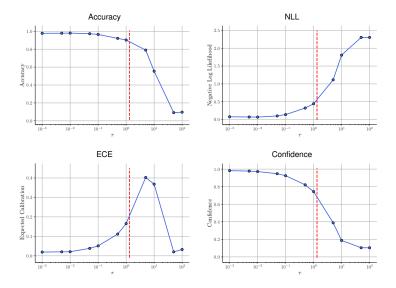


Figure: Effect of the temperature for a Linear BNN (one hidden layer, relU activations) trained on MNIST. No cooling $\eta_N = 1$ is indicated by a red line.

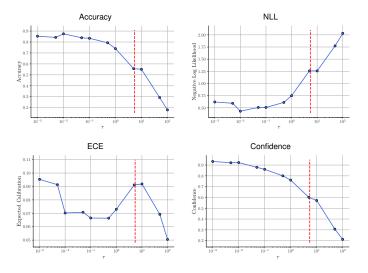


Figure: Effect of the temperature for a Resnet20 trained on CIFAR-10. No cooling $\eta_N = 1$ is indicated by a red line.

These experiments show that balancing the ELBO with the scaling $\eta_N = \tau p/N$ generalizes to much more complex architectures that a one hidden layer.

Conclusion

- We have identified that the ELBO should be tempered according to a temperature proportional to p/N, where p is the number of data points and N the number of parameters, when using product priors and posteriors
- With this choice, ELBO converges to a well-defined functional over the space of probability measures and one could analyze gradient descent dynamics through Wasserstein gradient flows
- Alternatively [Tran et al., 2020, Fortuin et al., 2021, Ober and Aitchison, 2021, Sun et al., 2019] have proposed the design of new priors which introduce correlation amongst the weights, however these models may be harder to train

Thank you! Questions?

References I

 Arbel, M., Korba, A., Salim, A., and Gretton, A. (2019).
 Maximum mean discrepancy gradient flow. arXiv preprint arXiv:1906.04370.

Ashukha, A., Lyzhov, A., Molchanov, D., and Vetrov, D. (2020).
 Pitfalls of in-domain uncertainty estimation and ensembling in deep learning.
 arXiv preprint arXiv:2002.06470.

Bhattacharya, A., Pati, D., and Yang, Y. (2019). Bayesian fractional posteriors. The Annals of Statistics, 47(1):39–66.

References II

 Blundell, C., Cornebise, J., Kavukcuoglu, K., and Wierstra, D. (2015).
 Weight uncertainty in neural network.
 In International Conference on Machine Learning, pages

1613–1622. PMLR. Chizat, L. and Bach, F. (2018).

On the global convergence of gradient descent for over-parameterized models using optimal transport. NIPS.

Coker, B., Pan, W., and Doshi-Velez, F. (2021).
 Wide mean-field variational bayesian neural networks ignore the data.

arXiv preprint arXiv:2106.07052.

References III

- Fortuin, V., Garriga-Alonso, A., Wenzel, F., Rätsch, G., Turner, R., van der Wilk, M., and Aitchison, L. (2021).
 Bayesian neural network priors revisited. arXiv preprint arXiv:2102.06571.
- Grünwald, P. (2012).
 - The safe bayesian.

In International Conference on Algorithmic Learning Theory, pages 169–183. Springer.

Grunwald, P., Steinke, T., and Zakynthinou, L. (2021).
 Pac-bayes, mac-bayes and conditional mutual information:
 Fast rate bounds that handle general vc classes.
 In Belkin, M. and Kpotufe, S., editors, *Proceedings of Thirty Fourth Conference on Learning Theory*, volume 134 of *Proceedings of Machine Learning Research*, pages 2217–2247. PMLR.

References IV

Grünwald, P. and Van Ommen, T. (2017). Inconsistency of bayesian inference for misspecified linear models, and a proposal for repairing it. *Bayesian Analysis*, 12(4):1069–1103.

Heide, R., Kirichenko, A., Grunwald, P., and Mehta, N. (2020).
 Safe-bayesian generalized linear regression.
 In International Conference on Artificial Intelligence and Statistics, pages 2623–2633. PMLR.

Izmailov, P., Vikram, S., Hoffman, M. D., and Wilson, A. G. (2021).

What are bayesian neural network posteriors really like? *International Conference on Machine Learning.*

References V

 Khan, M., Nielsen, D., Tangkaratt, V., Lin, W., Gal, Y., and Srivastava, A. (2018).
 Fast and scalable bayesian deep learning by weight-perturbation in adam.
 In International Conference on Machine Learning, pages 2611–2620. PMLR.

Mei, S., Montanari, A., and Nguyen, P.-M. (2018a).
 A mean field view of the landscape of two-layer neural networks.

Proceedings of the National Academy of Sciences, 115(33):E7665–E7671.

References VI

Mei, S., Montanari, A., and Nguyen, P.-M. (2018b).
 A mean field view of the landscape of two-layer neural networks.

Proceedings of the National Academy of Sciences, 115(33):E7665–E7671.

- Ober, S. W. and Aitchison, L. (2021).
 Global inducing point variational posteriors for bayesian neural networks and deep gaussian processes.
 In International Conference on Machine Learning, pages 8248–8259. PMLR.
- Osawa, K., Swaroop, S., Jain, A., Eschenhagen, R., Turner, R. E., Yokota, R., and Khan, M. E. (2019).
 Practical deep learning with bayesian principles. arXiv preprint arXiv:1906.02506.

References VII

Rotskoff, G., Jelassi, S., Bruna, J., and Vanden-Eijnden, E. (2019).
 Global convergence of neuron birth-death dynamics.
 In *ICML*.

Sun, S., Zhang, G., Shi, J., and Grosse, R. (2019).
 Functional variational bayesian neural networks.
 In International Conference on Learning Representations.

Tran, B.-H., Rossi, S., Milios, D., and Filippone, M. (2020). All you need is a good functional prior for bayesian deep learning.

arXiv preprint arXiv:2011.12829.

References VIII

 Wenzel, F., Roth, K., Veeling, B. S., Swiatkowski, J., Tran, L., Mandt, S., Snoek, J., Salimans, T., Jenatton, R., and Nowozin, S. (2020).
 How good is the bayes posterior in deep neural networks really?

International conference on machine learning.

Wilson, A. G. and Izmailov, P. (2020). Bayesian deep learning and a probabilistic perspective of generalization.

arXiv preprint arXiv:2002.08791.

 Zhang, G., Sun, S., Duvenaud, D., and Grosse, R. (2018). Noisy natural gradient as variational inference. In *International Conference on Machine Learning*, pages 5852–5861. PMLR.